



Signal Processing

Lecture 16: Dimensionality Reduction

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Why Dimensionality Reduction?

We often observe data in a *high-dimensional* space:

$$\mathbf{x}_i \in \mathbb{R}^D, \quad i = 1, \dots, N,$$

where D can be large (pixels, sensors, features, embeddings, . . .).

But the phenomenon that generated the data may be much simpler:

$$\mathbf{x}_i \approx g(\mathbf{z}_i), \quad \mathbf{z}_i \in \mathbb{R}^d, \quad d \ll D.$$

We want a representation

$$\mathbf{z}_i = g(\mathbf{x}_i) \in \mathbb{R}^d \quad (d \ll D)$$

that preserves the *important* structure of the data.

Typical goals:

- **Compression:** store/transmit \mathbf{x}_i using fewer numbers.
- **Visualization:** understand structure by plotting $d = 2$ or $d = 3$.
- **Denoising:** remove directions that are mostly noise.
- **Learning efficiency:** reduce computation and sample complexity.

As D grows, geometric intuition changes:

- Data become *sparse*: a fixed number of samples N covers a vanishing fraction of \mathbb{R}^D .
- Distances can become less informative (many points look “equally far”).
- Estimating quantities like densities or covariances becomes harder without much larger N .

In many problems, we pay a price in high dimensions:

more features \Rightarrow more parameters to estimate \Rightarrow more data needed.

High-dimensional features are often *correlated* or *redundant*.

Example idea:

$$\mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_D \end{bmatrix} \quad \text{with} \quad x_2 \approx 2x_1, \quad x_5 \approx x_1 + x_3, \quad \dots$$

So although $\mathbf{x} \in \mathbb{R}^D$, the data may concentrate near a much lower-dimensional set (e.g., a line, plane, or curved manifold).

Even if observations live in \mathbb{R}^D , the data may lie (approximately) in a lower-dimensional subspace:

$$\mathbf{x}_i \approx \boldsymbol{\mu} + \mathbf{W}\mathbf{z}_i, \quad \mathbf{W} \in \mathbb{R}^{D \times d}, \quad \mathbf{z}_i \in \mathbb{R}^d, \quad d \ll D.$$

Interpretation:

- $\boldsymbol{\mu}$ is the mean (translation).
- Columns of \mathbf{W} span a d -dimensional subspace capturing most variation.
- Remaining directions contain little signal (often noise).

Formalizing “keep what matters”

A common principle: preserve information measured by *variance* (or energy).

If \mathbf{P} is a projection onto a d -dimensional subspace, we want

$$\|\mathbf{x}_i - \mathbf{P}\mathbf{x}_i\|^2 \text{ small for most } i.$$

Equivalently, we want $\mathbf{P}\mathbf{x}_i$ to retain as much “spread” of the data as possible.

In other words, we want to:

- choose directions of **maximum variance**,
- or equivalently, minimize **reconstruction error**.

Let

$$\mathbf{A} = \begin{bmatrix} a & b \\ b & c \end{bmatrix} \quad (\text{symmetric}).$$

It defines a linear transformation $\mathbf{x} \mapsto \mathbf{Ax}$.

Key geometric fact (symmetric case):

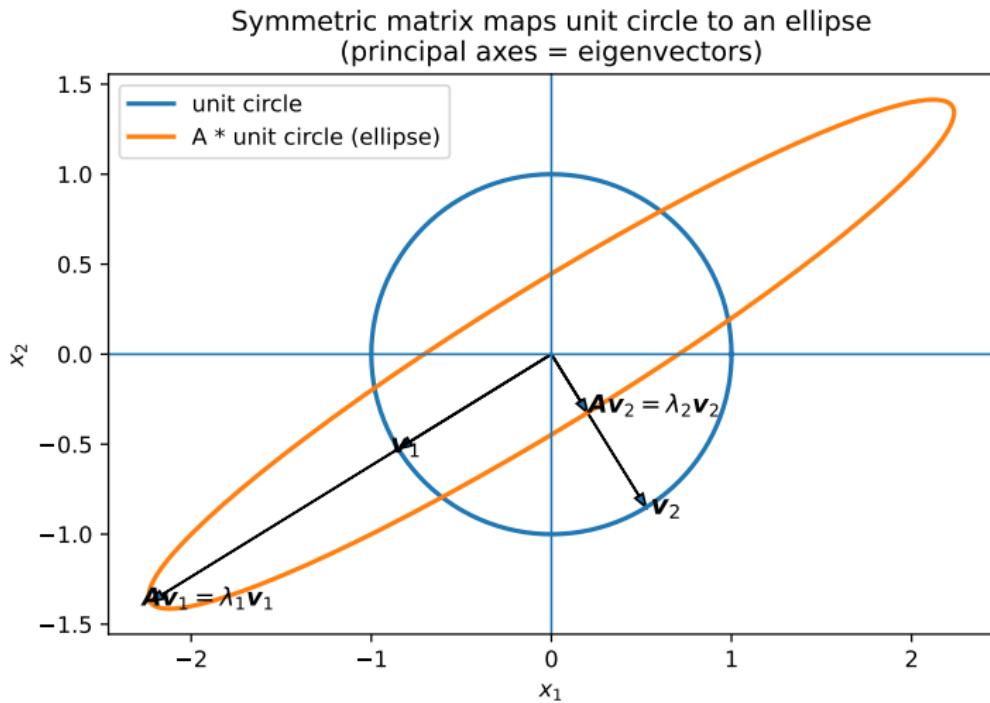
- \mathbf{A} maps the **unit circle** to an **ellipse**.
- The ellipse's principal axes are **orthogonal**.
- Those principal axes directions are the **eigenvectors** of \mathbf{A} .

Eigenpairs satisfy:

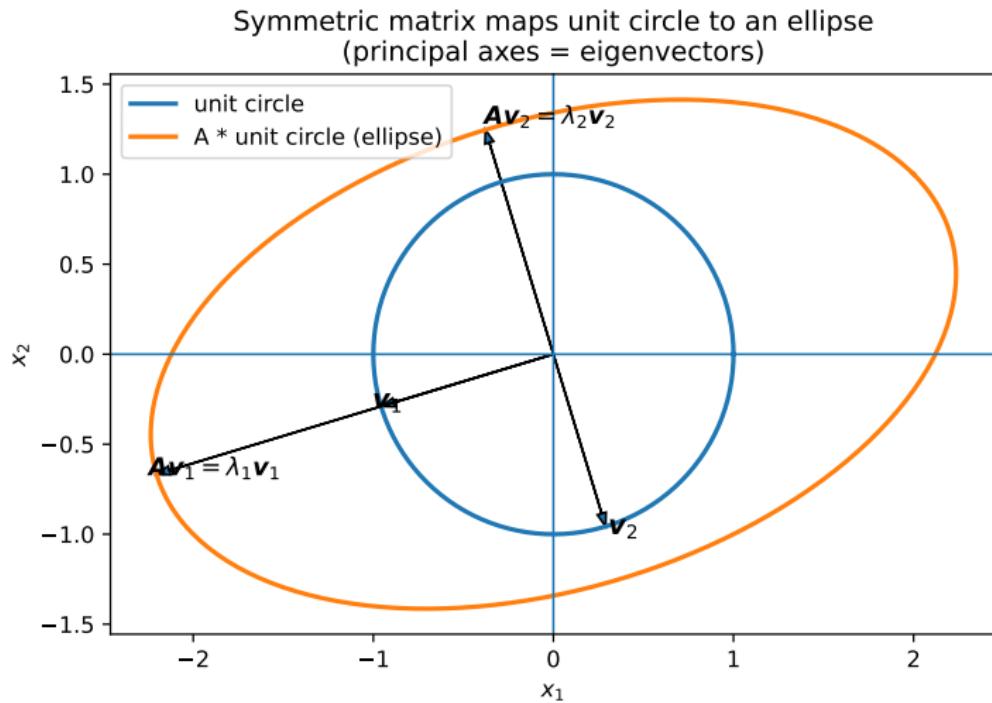
$$\mathbf{A}\mathbf{v}_k = \lambda_k \mathbf{v}_k.$$

Along direction \mathbf{v}_k , vectors are scaled by λ_k (no rotation).

Visualizing the transformation



Visualizing the transformation (2)



Unit circle \rightarrow ellipse (principal axes = eigenvectors)

Consider all unit vectors:

$$\mathcal{C} = \{\mathbf{x} \in \mathbb{R}^2 : \|\mathbf{x}\| = 1\}.$$

Their images under \mathbf{A} form:

$$\mathcal{E} = \{\mathbf{Ax} : \mathbf{x} \in \mathcal{C}\}.$$

For symmetric \mathbf{A} , \mathcal{E} is an ellipse.

If $\mathbf{A} = \mathbf{Q}\Lambda\mathbf{Q}^\top$ with

$$\Lambda = \text{diag}(\lambda_1, \lambda_2), \quad \mathbf{Q} = [\mathbf{v}_1 \ \mathbf{v}_2], \quad \mathbf{Q}^\top \mathbf{Q} = \mathbf{I},$$

then:

- the ellipse axes point along $\mathbf{v}_1, \mathbf{v}_2$,
- the axis “length scales” are $|\lambda_1|$ and $|\lambda_2|$ (applied to unit vectors).

Rayleigh quotient: “which direction stretches the most?”

For any nonzero \mathbf{x} define the Rayleigh quotient:

$$R(\mathbf{x}) = \frac{\mathbf{x}^\top \mathbf{A} \mathbf{x}}{\mathbf{x}^\top \mathbf{x}}.$$

If $\|\mathbf{x}\| = 1$, then $R(\mathbf{x}) = \mathbf{x}^\top \mathbf{A} \mathbf{x}$.

For symmetric \mathbf{A} :

$$\max_{\|\mathbf{x}\|=1} \mathbf{x}^\top \mathbf{A} \mathbf{x} = \lambda_{\max} \quad \text{and is achieved at } \mathbf{x} = \mathbf{v}_{\max}.$$

$$\min_{\|\mathbf{x}\|=1} \mathbf{x}^\top \mathbf{A} \mathbf{x} = \lambda_{\min} \quad \text{and is achieved at } \mathbf{x} = \mathbf{v}_{\min}.$$

What changes for a non-symmetric matrix?

Let

$$\mathbf{A} \in \mathbb{R}^{2 \times 2}, \quad \mathbf{A} \neq \mathbf{A}^T.$$

Then $\mathbf{x} \mapsto \mathbf{A}\mathbf{x}$ can include **shear** and **stretching-like** effects.

Key differences from the symmetric case:

- Eigenvalues may be **complex** (no real invariant directions).
- Real eigenvectors (if they exist) need **not be orthogonal**.
- The image of the unit circle is still an ellipse, but its principal axes are **not** generally given by eigenvectors.

Singular Value Decomposition (SVD)

For any matrix

$$\mathbf{A} \in \mathbb{R}^{m \times n},$$

there exists a decomposition

$$\mathbf{A} = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T,$$

where:

- $\mathbf{U} \in \mathbb{R}^{m \times m}$ has orthonormal columns,
- $\mathbf{V} \in \mathbb{R}^{n \times n}$ has orthonormal columns,
- $\mathbf{\Sigma} \in \mathbb{R}^{m \times n}$ is diagonal with

$$\sigma_1 \geq \sigma_2 \geq \dots \geq 0.$$

Geometric interpretation of $\mathbf{x} \mapsto \mathbf{A}\mathbf{x}$:

$\mathbf{x} \xrightarrow{\mathbf{V}^T}$ rotate/reflect $\xrightarrow{\mathbf{\Sigma}}$ axis-aligned scaling $\xrightarrow{\mathbf{U}}$ rotate/reflect.

Consider the unit circle $\{x : \|x\| = 1\}$ in \mathbb{R}^2 .

Under $x \mapsto Ax$:

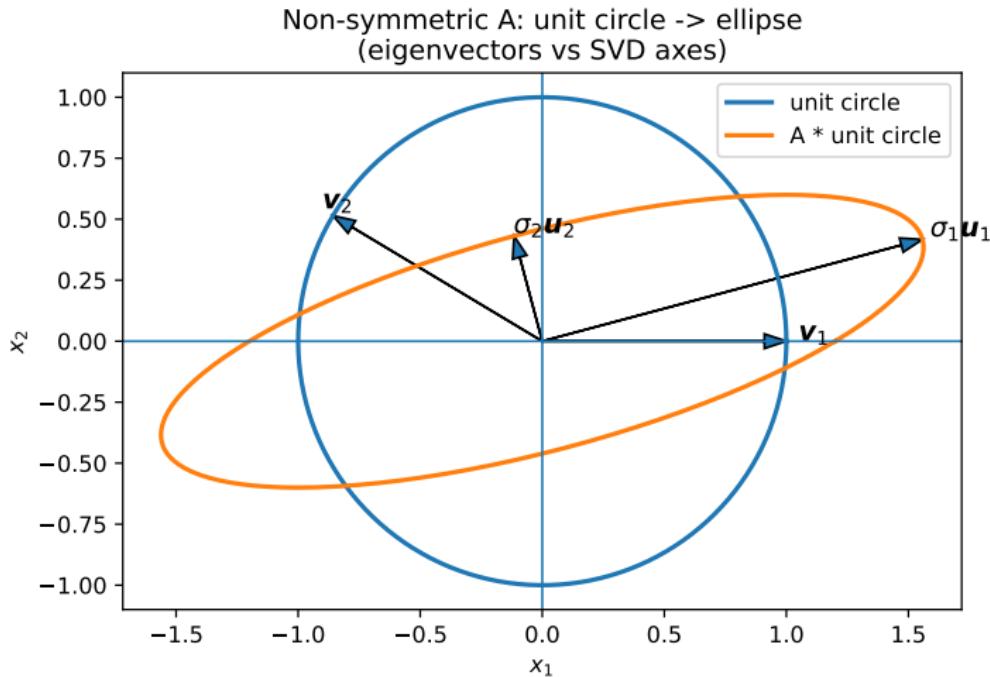
- columns of V give **input directions** that map to the ellipse axes,
- columns of U give the **output directions** of those axes,
- singular values σ_k are the **lengths** of the ellipse semi-axes.

Unlike eigenvectors:

- SVD always exists (for any matrix),
- directions are always orthogonal,
- geometry is always real and well-defined.

Key takeaway: eigenvectors explain symmetric matrices well;
SVD explains *all* linear maps.

Visualizing SVD



Principal Component Analysis (PCA)

Given centered data $\tilde{\mathbf{x}}_i = \mathbf{x}_i - \bar{\mathbf{x}}$ and covariance

$$\boldsymbol{\Sigma} = \frac{1}{N} \sum_{i=1}^N \tilde{\mathbf{x}}_i \tilde{\mathbf{x}}_i^\top,$$

PCA finds an orthonormal basis $\mathbf{W}_d = [\mathbf{w}_1 \ \cdots \ \mathbf{w}_d]$ that best captures the data.

Variance-maximizing view:

$$\mathbf{w}_1 = \arg \max_{\|\mathbf{w}\|=1} \mathbf{w}^\top \boldsymbol{\Sigma} \mathbf{w}, \quad \mathbf{w}_k = \arg \max_{\|\mathbf{w}\|=1, \ \mathbf{w} \perp \mathbf{w}_1, \dots, \mathbf{w}_{k-1}} \mathbf{w}^\top \boldsymbol{\Sigma} \mathbf{w}.$$

Equivalent reconstruction view:

$$\mathbf{W}_d = \arg \min_{\mathbf{W}^\top \mathbf{W} = I} \sum_{i=1}^N \left\| \tilde{\mathbf{x}}_i - \mathbf{W} \mathbf{W}^\top \tilde{\mathbf{x}}_i \right\|^2.$$

Solution: \mathbf{w}_k are the eigenvectors of $\boldsymbol{\Sigma}$, with eigenvalues

$$\boldsymbol{\Sigma} \mathbf{w}_k = \lambda_k \mathbf{w}_k, \quad \lambda_1 \geq \lambda_2 \geq \cdots \geq 0,$$

and the reduced coordinates are

$$\mathbf{z}_i = \mathbf{W}_d^\top \tilde{\mathbf{x}}_i.$$

PCA Example: 3D points on a plane

We observe points in 3D:

$$\mathbf{x}_i \in \mathbb{R}^3, \quad i = 1, \dots, N.$$

But the data are *approximately planar* (intrinsic dimension ≈ 2):

$$\mathbf{x}_i \approx \boldsymbol{\mu} + u_i \mathbf{p}_1 + v_i \mathbf{p}_2 + \varepsilon_i \mathbf{n},$$

where $\mathbf{p}_1, \mathbf{p}_2$ span an (unknown) plane, \mathbf{n} is its normal, and $|\varepsilon_i|$ is small.

Within that plane, the points form an “S” shape (nonlinear structure), so:

- dimension $3 \rightarrow 2$ should work well,
- dimension $3 \rightarrow 1$ will generally *lose* structure.

PCA goal (variance-maximizing view)

First, center the data:

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{i=1}^N \mathbf{x}_i, \quad \tilde{\mathbf{x}}_i = \mathbf{x}_i - \bar{\mathbf{x}}.$$

Define the sample covariance:

$$\boldsymbol{\Sigma} = \frac{1}{N} \sum_{i=1}^N \tilde{\mathbf{x}}_i \tilde{\mathbf{x}}_i^\top \in \mathbb{R}^{3 \times 3}.$$

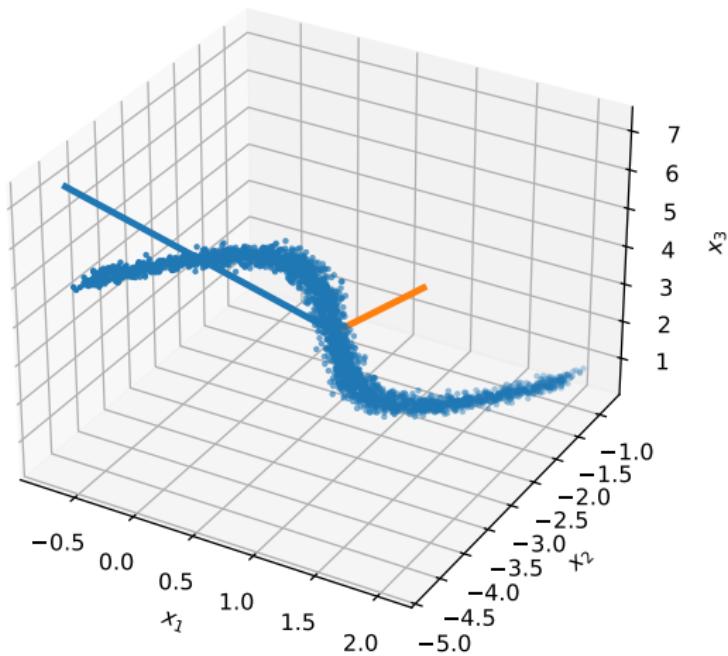
The first principal component solves

$$\max_{\|\mathbf{w}\|=1} \text{Var}(\mathbf{w}^\top \tilde{\mathbf{x}}) = \max_{\|\mathbf{w}\|=1} \mathbf{w}^\top \boldsymbol{\Sigma} \mathbf{w}.$$

Thus \mathbf{w}_1 is the top eigenvector of $\boldsymbol{\Sigma}$, and the variance along it is λ_1 .

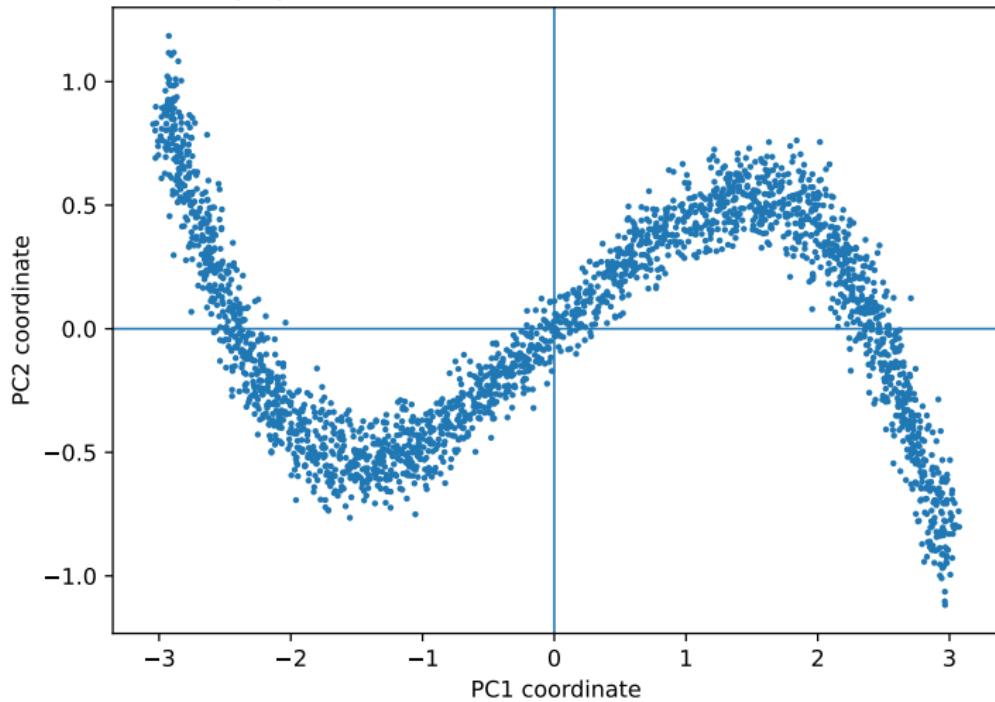
Visualizing PCA

3D data: 'S' on a rotated plane + PCA axes



Visualizing PCA (2)

PCA projection to 2D (PC1 vs PC2) — should show the 'S'



Why PCA discovers the plane (in our example)

Because the points are (approximately) on a plane:

λ_1 large, λ_2 large, λ_3 small.

Geometrically:

- $\mathbf{w}_1, \mathbf{w}_2$ span the best-fitting plane (maximum retained variance),
- \mathbf{w}_3 approximates the plane normal (minimum variance direction).

We reduce dimension by projection:

$$\mathbf{W}_2 = [\mathbf{w}_1 \ \mathbf{w}_2] \in \mathbb{R}^{3 \times 2}, \quad \mathbf{z}_i = \mathbf{W}_2^\top \tilde{\mathbf{x}}_i \in \mathbb{R}^2.$$

The 2D coordinates \mathbf{z}_i should reveal the “S” clearly.

From reduced coordinates:

$$\hat{\mathbf{x}}_i = \bar{\mathbf{x}} + \mathbf{W}_2 \mathbf{z}_i$$

is the orthogonal projection of \mathbf{x}_i onto the PCA plane.

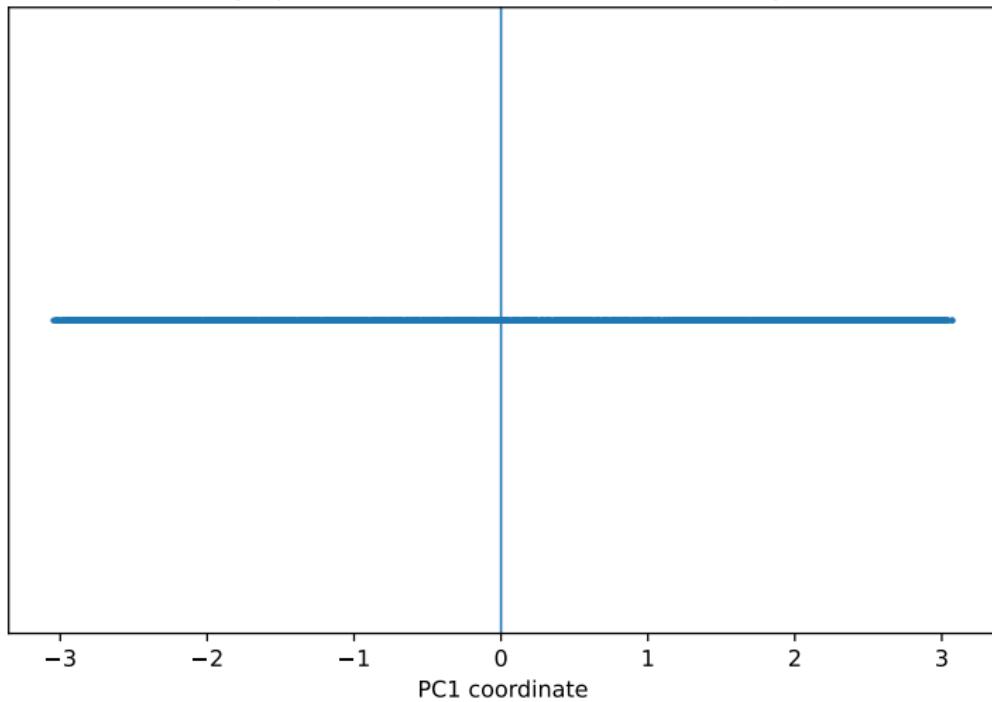
Reconstruction error (per point):

$$\|\mathbf{x}_i - \hat{\mathbf{x}}_i\|^2 = \|\tilde{\mathbf{x}}_i - \mathbf{W}_2 \mathbf{W}_2^\top \tilde{\mathbf{x}}_i\|^2.$$

If we keep only 1 component ($\mathbf{W}_1 = \mathbf{w}_1$), the 1D projection can *collapse* parts of the “S” onto each other (structure loss).

Visualizing PCA (3)

PCA projection to 1D (PC1) — structure collapses



PCA via SVD (same geometry, two viewpoints)

Stack centered samples as rows of a data matrix:

$$\tilde{\mathbf{X}} = \begin{bmatrix} \tilde{\mathbf{x}}_1^\top \\ \tilde{\mathbf{x}}_2^\top \\ \vdots \\ \tilde{\mathbf{x}}_N^\top \end{bmatrix} \in \mathbb{R}^{N \times D}, \quad \tilde{\mathbf{x}}_i = \mathbf{x}_i - \bar{\mathbf{x}}.$$

Compute the SVD:

$$\tilde{\mathbf{X}} = \mathbf{U} \boldsymbol{\Sigma} \mathbf{V}^\top, \quad \boldsymbol{\Sigma} = \text{diag}(\sigma_1, \sigma_2, \dots).$$

Then the sample covariance is

$$\boldsymbol{\Sigma}_{\text{data}} = \frac{1}{N} \tilde{\mathbf{X}}^\top \tilde{\mathbf{X}} = \mathbf{V} \left(\frac{\boldsymbol{\Sigma}^2}{N} \right) \mathbf{V}^\top.$$

So:

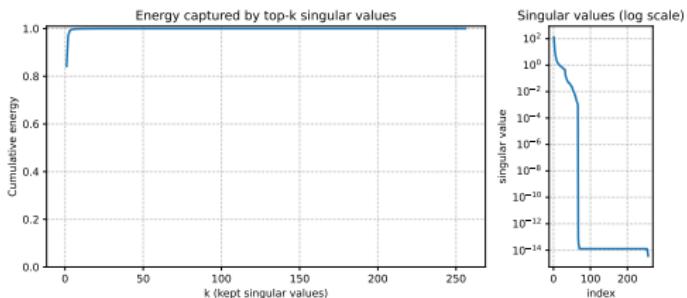
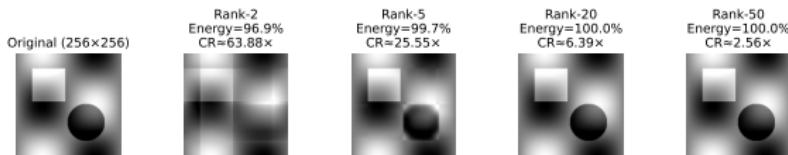
- **PCA directions** (principal axes) are columns of \mathbf{V} .
- **PCA eigenvalues** are $\lambda_k = \sigma_k^2/N$.
- **Reduced coordinates (scores):**

$$\mathbf{z}_i = \mathbf{V}_d^\top \tilde{\mathbf{x}}_i \quad \Leftrightarrow \quad \tilde{\mathbf{X}} \mathbf{V}_d = \mathbf{U}_d \boldsymbol{\Sigma}_d.$$

SVD for Image Compression (Grayscale)

- Treat an image as a matrix $\mathbf{X} \in \mathbb{R}^{m \times n}$ (pixel intensities).
- Compute SVD: $\mathbf{X} = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T$.
- Keep top- k singular values/vectors for compression:

$$\mathbf{X}_k = \mathbf{U}_{[:,1:k]} \mathbf{\Sigma}_{1:k,1:k} \mathbf{V}_{[:,1:k]}^T.$$



Thank you

- Any Questions?

- Office Hours:

- Tue & Thu (09:00-11:00)

- 24/7 by email (costashatz@upatras.gr, subject: *ECE_SP_AM*)

- Material and Announcements



Laboratory of Automation & Robotics